Comparison analysis of two alternative hedging methods for CDO tranches

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Introduction and motivation

- Recent financial turmoil has deeply affected the market of structured credit derivatives
- CDS index products are still liquid but ...
- Investors have more incentive to risk-manage their trading books
- However, standard hedging methods have revealed serious drawbacks during the crisis
 - focus on the computation of spread sensitivities (credit deltas)
 - within a static Gaussian copula model
 - does not rely on a sound theory of replication
 - negative deltas may occur in a steep base correlation market: Morgan and Mortensen (2007)

Introduction and motivation

- We consider discrete-time hedging of index CDO tranches using
 - the underlying CDS index
 - the risk-free asset
- Hedging consists in taking complementary positions in the index and in the risk-free asset in order to minimize the overall evolution of market prices.
 - These positions need to be regularly updated over time
- Aim of the presentation: performance analysis of two alternative hedging strategies
 - $\Delta^{lo}:$ delta of the tranche within a Markovian contagion model
 - Δ^{li} : delta of the tranche within a Gaussian copula model

In the literature

- Crépey (2004) performs a similar analysis for the equity market
 - Comparison of hedging performance of equity options using two alternative deltas: Black-Scholes implied delta and local volatility delta
 - He exhibits two market directions: (slow/fast) and (rallies/sell-offs)
 - Negatively skew market: local volatility delta provides a better hedge than implied delta during slow rallies or fast sell-offs and a worse hedge during fast rallies and slow sell-offs.
- Analogies with the credit market are not so obvious
 - Interaction between default risk and spread risk, large dimension of the portfolio, recovery rate uncertainty

In the literature

- Laurent, Cousin and Fermanian (2007) study the hedging of index CDO tranche in a Markovian contagion model using the CDS index
 - When simultaneous defaults are precluded, the CDO tranche market is complete
 - Computation of dynamic hedging strategies along the nodes of a binomial tree
 - Δ^{lo} seems to be smaller than Δ^{li} for equity tranches and higher for more senior tranches when the two models are calibrated to the same market data.
 - Do not study in details the performance of these two hedging strategies

In the literature

- Cont and Kan (2008) perform an empirical comparison of various hedging strategies for index CDO tranches
 - three different notions of deltas: spread-delta, jump-to-default delta, quadratic risk-minimization delta
 - deltas computed in various models calibrated to the same set of market data
 - Backtest the strategies before and during the crisis
- Main conclusions:
 - spread-deltas are very similar across models calibrated to the same data set
 - jump-to-default ratios are significatively different across models (substantial model risk)
 - spread-deltas hedges are preferred before the crisis and not preferred during the crisis
- But study of performance only for a single (market) trajectory
- Do not address the issue of hedging with individual CDS

CDS Index and standardized CDO tranches

- Slice the credit portfolio into different risk levels or CDO tranches
- ex: CDO tranche on standardized Index such as CDX North America or Itraxx Europe



Notations

- Credit portfolio with n reference entities
- τ_1, \ldots, τ_n : default times
- R: homogeneous and constant recovery rate at default (R = 40% typically)
- Number of defaults process:

$$N_t = \sum_{i=1}^n \mathbb{1}_{\{\tau_i \le t\}}$$

• CDO tranche cash-flows are driven by the aggregate loss process normalized to unity:

$$L_t = \frac{1}{n}(1-R)N_t$$

• Cash-flows only depends on $\phi(L_t)$, $0 \le t \le T$

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Assumptions

- By simplicity, we consider zero interest rate r = 0
- Stylized products with simplified cash-flows:
 - Cash-flows of the index and associated CDO tranches only depend on $\phi(L_T)$
 - $\bullet\,$ Protection or default payment only occur at maturity T
 - The premium leg is paid upfront
- Given a risk-neutral probability \mathbb{P} and a fitration $\mathbb{F} = (\mathcal{F}_t)_{t \in [0,T]}$
 - The time t price of a derivative with a $\mathcal{F}_T\text{-measurable}$ and bounded payoff $\xi=\phi(L_T)$ is:

$$\mathbb{E}[\xi \mid \mathcal{F}_t]$$

Products and models under scrutiny Comparison of two hedging strategies Future research

CDS Index and standardized CDO tranches Models under consideration

CDS Index and standardized CDO tranches

• Loss on CDO tranche [a, b]: $L_T^{[a, b]}$ has a call spread payoff with respect to L_T :



- Loss on CDO tranche [a, b]: $L_T^{[a,b]} = (L_T - a)^+ - (L_T - b)^+ = \phi^{[a,b]}(N_T)$
 - equity tranches: a = 0% and $L_t^{[0,b]} = \min(L_T, b)$
 - senior tranches: b = 100% and $L_t^{[a,1]} = (L_T a)^+$
 - CDS index associated with a [0%,100%] CDO tranche

CDS Index and standardized CDO tranches

- As the premium leg is paid upfront, the analysis is focused on the **protection leg**
- The time-t cum-dividend price of a stylized CDO tranche [a, b] (protection leg) is referred to as:

$$\Pi_t = \mathbb{E}\left[\phi^{[a,b]}(N_T) \mid \mathcal{F}_t\right]$$

• The time-*t* cum-dividend price of the stylized underlying index (protection leg) is referred to as:

$$P_t = \mathbb{E}\left[\phi^{[0,1]}(N_T) \mid \mathcal{F}_t\right]$$

Homogeneous one factor Gaussian copula model

 $\bullet\,$ Also referred to as the $Li\ model$

•
$$V_i = \rho V + \sqrt{1 - \rho^2} \bar{V_i}, \ i = 1 \dots n$$
: latent variables

- $V, \bar{V}_i, i = 1 \dots n$: independent Gaussian random variables
- Default times defined by: $\tau_i = F_i^{-1}(\Phi(V_i)), i = 1 \dots n$
 - $F_1 = \ldots = F_n = F$: cdf of $\tau_i, i = 1, \ldots, n$
 - Φ : cdf of V_i
- Conditional default probability

$$p_t(V) = \mathbb{P}(\tau_i \le t \mid V) = \Phi\left(\frac{\Phi^{-1}(F(t)) - \rho V}{\sqrt{1 - \rho^2}}\right)$$

• Loss distribution is merely a binomial mixture:

$$\mathbb{P}(N_t = k) = \binom{n}{k} \int p_t(x)^k (1 - p_t(x))^{n-k} \nu(x) dx, \quad k = 0, \dots, n$$

Homogeneous one factor Gaussian copula model

- $\bullet\,$ At each time t, the model parameters ρ_t and F_t are calibrated on market spreads
- F_t is inferred from the term structure of index spreads at time t
 - ${\ensuremath{\, \rm o}}$ Index spread curve assumed to be flat and equal to S_t

$$F_t(s) = \mathbb{P}(\tau_i \le t) = 1 - \exp\left(-\frac{S_t}{1-R}(s-t)\right), \ s \ge t$$

- One dependence parameter ρ_t^b associated with each base tranche [0,b], b= 3%, 6%, 9%, 12%, 22% (iTraxx)
 - $\Pi^{ma}_t(T,a,b)$: market price of CDO tranche [a,b], maturity T
 - $\Pi^{li}(T, a, b; t, S_t, \rho_t)$: price of CDO tranche [a, b] in the Li model
 - Base correlation ρ_t^b is such that:

$$\Pi^{li}(T, 0, b; t, S_t, \rho_t^b) = \Pi_t^{ma}(T, 0, b)$$

Homogeneous one factor Gaussian copula model

 $\bullet\,$ Monotonic base tranche and senior tranche prices with respect to ρ in the Li model

$$\frac{\partial \Pi^{li}(T,0,b;t,S_t,\rho)}{\partial \rho} \le 0, \quad \frac{\partial \Pi^{li}(T,a,1;t,S_t,\rho)}{\partial \rho} \ge 0, \quad \forall a,b \in [0,1]$$

- Given the existence of base correlations $\rho_t^b,$ these parameters are unique
- CDO tranche market typically reflects steep base correlation curves:



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Homogeneous Markovian contagion model

- Also referred to as the local intensity model
- Dynamic model where the cumulative default intensity only depends on number of defaults
- N_t is a continuous-time Markov chain with generator matrix:

$$\Lambda(t) = \begin{pmatrix} -\lambda(t,0) & \lambda(t,0) & 0 & & 0\\ 0 & -\lambda(t,1) & \lambda(t,1) & & 0\\ & & \ddots & \ddots & \\ 0 & & & -\lambda(t,n-1) & \lambda(t,n-1)\\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}$$

• $\Pi^{lo}(t,T) = \mathbb{E}\left[\Phi(N_T) \mid \mathcal{F}_t\right] = \mathbb{E}\left[\Phi(N_T) \mid N_t\right]$

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CDS Index and standardized CDO tranches Models under consideration

Homogeneous Markovian contagion model

- Vector of prices $\Pi^{lo}(t,T) = \left(\Pi_1^{lo}(t,T),\ldots,\Pi_n^{lo}(t,T)\right)^\top$
 - where $\Pi_i^{lo}(t,T) = \mathbb{E}\left[\Phi(N_T) \mid N_t = i\right]$, $i = 1, \dots, n$
- can be related to the vector of terminal payoffs $C = \left(\Phi(0), \ldots, \Phi(n)
 ight)^{ op}$
- using the backward Kolmogorov equation:

$$\begin{cases} \frac{\partial \Pi^{lo}(t,T)}{\partial t} = -\Lambda(t)\Pi^{lo}(t,T) \\ \Pi^{lo}(T,T) = C \end{cases}$$

• When the intensities are time-homogeneous, i.e, $\Lambda(t) = \Lambda$ then:

$$\Pi^{lo}(t,T) = \exp\left((T-t)\Lambda\right)C$$

• Approach puts in practice by van der Voort (2006), Herbersson (2007), Laurent, Cousin and Fermanian (2007), Arnsdorf and Halperin (2007), Lopatin and Misirpashaev (2007), Cont and Minca (2008)

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Delta-hedging in discrete time

- An investor enters a sell-protection position on a CDO tranche [a, b]
- He wants to cover his position until an hedging horizon: $T_1 \leq T$
- **Delta-hedging** the tranche consists in rebalancing a complementary position in a portfolio including the underlying index and the risk-free asset
 - at every point in time of a subdivision $0 = t_0 \le t_1 \le \cdots \le t_p = T_1$
 - index position determined in order to minimize the overall exposure

Delta-hedging in discrete time

• The profit-and-loss (P&L) trajectory $e = (e_{t_k})_{0 \le k \le p}$ is obtained by adding up the P&L increments:

$$\delta_k e = -\delta_k \Pi + \Delta_{t_k} \delta_k P$$

- $\delta_k \Pi = \Pi_{t_{k+1}} \Pi_{t_k} :$ increments of the tranche market price in $(t_k, t_{k+1}]$
- $\delta_k P = P_{t_{k+1}} P_{t_k}$: increments of the index market price in $(t_k, t_{k+1}]$
- Δ_{t_k} : number of units of index contract in the hedging portfolio over the time interval $(t_k, t_{k+1}]$
- Aim is to compare the P&L trajectory e obtained using two strategies:
 - $\Delta=\Delta^{lo}\!\!:$ delta of the tranche in a Markovian contagion model
 - $\Delta=\Delta^{li}:$ delta of the tranche in a Gaussian copula model

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Delta-hedging in discrete time

• $\Delta^{lo}:$ jump-to-default in the local intensity model

$$\Delta_t^{lo} = \frac{\Pi_{i+1}^{lo}(t) - \Pi_i^{lo}(t)}{P_{i+1}^{lo}(t) - P_i^{lo}(t)}$$

• where
$$\Pi_i^{lo}(t) = \mathbb{E}\left[\Phi^{[a,b]}(N_T) \mid N_t = i\right]$$

• and $P_i^{lo}(t) = \mathbb{E}\left[\Phi^{[0,1]}(N_T) \mid N_t = i\right]$

• Δ^{li} : spread-delta (sticky strike rule) in the Li model

$$\Delta_t^{li} = \frac{\Pi^{li}(t, S_t + \varepsilon, \rho_t) - \Pi^{li}(t, S_t, \rho_t)}{P^{li}(t, S_t + \varepsilon) - P^{li}(t, S_t)}$$

• ε is typically equal to some few basis points

• The two models are calibrated on the same set of market spreads at every time t_k , $k=0,\ldots,p-1$

- Methodology similar to Hull and Suo (2000) and Crépey (2004)
- Theoretical market given as a fixed Markovian contagion model:
 - Market prices are such that:

 $\Pi_t = \Pi^{lo}(t, N_t), \ P_t = P^{lo}(t, N_t), \ S_t = S^{lo}(t, N_t), \ \rho_t = \rho^{lo}(t, N_t)$

- Given $N_t = i$, $\Pi_t = \Pi_i^{lo}(t)$, $P_t = P_i^{lo}(t)$, $S_t = S_i^{lo}(t)$, $\rho_t = \rho_i^{lo}(t)$
- As the CDO tranche market is complete in homogeneous Markovian contagion model, Δ^{lo} is the perfect continuous-time hedging strategy
- But we consider hedging in discrete time

 $\bullet\,$ Simulation of \bar{N} default trajectories in the local intensity model

• Simulation of
$$\left(au^{(1)},\ldots, au^{(n)}
ight)_{j}$$
, $j=1,\ldots,ar{N}$

• Without loss of generality, we focus the hedging analysis on a single period $(t_k, t_{k+1}]$

•
$$\Delta_{t_k}^{lo}$$
 is preferred to $\Delta_{t_k}^{li}$ on the period $(t_k, t_{k+1}]$ if

$$\mathsf{Var}\,(\delta_k e^{lo}) < \mathsf{Var}\,(\delta_k e^{li})$$

• where $\delta_k e^{lo}$ is the P&L increment in $(t_k, t_{k+1}]$ using $\Delta_{t_k}^{lo}$ • where $\delta_k e^{li}$ is the P&L increment in $(t_k, t_{k+1}]$ using $\Delta_{t_k}^{li}$.

- As in Crépey (2004) we distinguish two market directions (slow/fast) and (rallies/sell-offs)
- Regarding the period $(t_k, t_{k+1}]$, a market trajectory is said to be
 - fast: a default is observed on the period $(t_k, t_{k+1}]$
 - slow: no default is observed on the period $(t_k, t_{k+1}]$
 - rallies: $\delta_k P \leq 0$ (decreasing index spread)
 - sell-offs: $\delta_k P \ge 0$ (increasing index spread)

• We consider the hedging of an **equity tranche** (Analysis is similar for a senior tranche)

Proposition

 $\delta_k e^{lo}$ is **positive** at **slow** market regimes and **negative** at **fast** market regimes

• Indeed, one can remark that:

$$\delta_k e^{lo} = -\delta_k \Pi + \Delta_{t_k}^{lo} \delta_k P = \int_{t_k}^{t_{k+1}} \left(\Delta_{t_k}^{lo} - \Delta_t^{lo} \right) dP_t$$

- Consider a small interval $(t_k, t_{k+1}]$
- no default (slow): $dP_t \simeq \delta_k P \leq 0$ (time decay effect)
- one default (fast): $dP_t \simeq \delta_k P \ge 0$ (cash-flow and contagion effect)

Definition of hedge ratios Analysis in a market governed by a Markovian contagion model

Analysis in a Markovian contagion model

• And Δ_t^{lo} is typically increasing in t:



•
$$(t,i) \to \Delta^{lo}(t,i) = \frac{\prod_{i=1}^{lo}(t) - \prod_{i=0}^{lo}(t)}{P_{i+1}^{lo}(t) - P_{i}^{lo}(t)}, \ 0 \le t \le 5, \ i = 0, \dots, 6$$

• Ordering of the two deltas for equity tranche

Proposition

- If $\rho_{i+1}^{lo}(t) \geq \rho_i^{lo}(t),$ then one may expect that $\Delta_t^{lo} \leq \Delta_t^{li}$
- If $\rho_{i+1}^{lo}(t) \leq \rho_{i}^{lo}(t),$ then one may expect that $\Delta_{t}^{lo} \geq \Delta_{t}^{li}$
- Indeed, by definition of the implied base correlation:

$$\begin{aligned} \Pi_{i+1}^{lo}(t) &- \Pi_{i}^{lo}(t) &= \Pi^{li}\left(t, S_{i+1}^{lo}(t), \rho_{i+1}^{lo}(t)\right) - \Pi^{li}\left(t, S_{i}^{lo}(t), \rho_{i}^{lo}(t)\right) \\ &= \Pi^{li}\left(t, S_{i+1}^{lo}(t), \rho_{i+1}^{lo}(t)\right) - \Pi^{li}\left(t, S_{i+1}^{lo}(t), \rho_{i}^{lo}(t)\right) \\ &+ \Pi^{li}\left(t, S_{i+1}^{lo}(t), \rho_{i}^{lo}(t)\right) - \Pi^{li}\left(t, S_{i}^{lo}(t), \rho_{i}^{lo}(t)\right) \end{aligned}$$

• But as $\partial_{\rho}\Pi^{li}(t, S, \rho) \leq 0$ for an equity tranche:

$$\Pi^{li}\left(t, S^{lo}_{i+1}(t), \rho^{lo}_{i+1}(t)\right) \leq \Pi^{li}\left(t, S^{lo}_{i+1}(t), \rho^{lo}_{i}(t)\right)$$

• Ordering of the two deltas (cont.)

Δ

• By definition of the local intensity delta:

$$\begin{split} {}^{lo}(t,i) &= \Delta^{lo}(t,i) = \frac{\Pi_{i+1}^{lo}(t) - \Pi_{i}^{lo}(t)}{P_{i+1}^{lo}(t) - P_{i}^{lo}(t)} \\ &\leq \frac{\Pi^{li}\left(t, S_{i+1}^{lo}(t), \rho_{i}^{lo}(t)\right) - \Pi^{li}\left(t, S_{i}^{lo}(t), \rho_{i}^{lo}(t)\right)}{P_{i+1}^{lo}(t) - P_{i}^{lo}(t)} \\ &= \frac{\Pi^{li}\left(t, S_{i+1}^{lo}(t), \rho_{i}^{lo}(t)\right) - \Pi^{li}\left(t, S_{i}^{lo}(t), \rho_{i}^{lo}(t)\right)}{P^{li}\left(t, S_{i+1}^{lo}(t)\right) - P^{li}\left(t, S_{i}^{lo}(t)\right)} \\ &\simeq \Delta_{t}^{li} \end{split}$$

• Comparison of P&L increments obtained using Δ^{lo} and Δ^{li}

$$\delta e^{li} = \delta e^{lo} + \left(\Delta^{li} - \Delta^{lo}\right) \delta P$$

- In a market where $\rho_{i+1}^{lo}(t) \ge \rho_i^{lo}(t)$ (steep base correlation market)
- [0% b] equity tranche:

Market regime	Rally	Sell-Off
Slow	$(\delta e^{li})^+ \le \delta e^{lo}$	
Fast		$\delta e^{lo} \le -(\delta e^{li})^-$

• [a - 100%] senior tranche:

Market regime	Rally	Sell-Off
Slow	$\delta e^{lo} \le -(\delta e^{li})^-$	
Fast		$(\delta e^{li})^+ \le \delta e^{lo}$

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 \bullet Comparison of P&L increments obtained using Δ^{lo} and Δ^{li}

$$\delta e^{li} = \delta e^{lo} + \left(\Delta^{li} - \Delta^{lo}\right) \delta P$$

- In a market where $\rho_{i+1}^{lo}(t) \leq \rho_i^{lo}(t)$ (flat base correlation market)
- [0% b] equity tranche:

Market regime	Rally	Sell-Off
Slow	$0 \le \delta e^{lo} \le \delta e^{li}$	
Fast		$\delta e^{li} \le \delta e^{lo} \le 0$

•
$$[a - 100\%]$$
 senior tranche:

Market regime	Rally	Sell-Off
Slow	$\delta e^{li} \le \delta e^{lo} \le 0$	
Fast		$0 \le \delta e^{lo} \le \delta e^{li}$

• Δ^{lo} provides a better hedge than Δ^{li}

Numerical results

- Sell-protection position in a [0-3%] equity tranche
- We numerically compare hedging performance of Δ^{li} and Δ^{lo} in two different markets:
 - One with a high increase of contagion (red)
 - One with a low increase of contagion (blue)



• The high contagion market features an increase of base correlation at the arrival of defaults:



$$\rho_{i+1}^{lo}(t) \ge \rho_i^{lo}(t)$$

- Histogram of P&L increments using Δ^{lo} (left) and Δ^{li} (right)
 - Hedging period: [0, 0.02] (one week)
 - $\bar{N} = 10000$ trajectories



- Histogram of P&L increments using Δ^{lo} (left) and Δ^{li} (right)
 - Hedging period: [0, 0.09] (one month)
 - $\bar{N} = 10000$ trajectories



- Histogram of P&L increments using Δ^{lo} (left) and Δ^{li} (right)
 - Hedging period: [0, 0.5] (one semester)
 - $\bar{N} = 10000$ trajectories



- Standard deviation of P&L increments function of the hedging horizon
 - $\bar{N}=10000$ simulations at each time step
 - $\sigma(\delta P\&L)$ using Δ^{lo} in blue
 - $\sigma(\delta P \& L)$ using Δ^{li} in red



• The low contagion market features a decrease of base correlation at the arrival of defaults:



 $\rho_{i+1}^{lo}(t) \leq \rho_i^{lo}(t)$

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Products and models under scrutiny Comparison of two hedging strategies Future research

Definition of hedge ratios Analysis in a market governed by a Markovian contagion model

- Histogram of P&L increments using Δ^{lo} (left) and Δ^{li} (right)
 - Hedging period: [0, 0.02] (one week)
 - $\bar{N} = 10000$ trajectories



Products and models under scrutiny Comparison of two hedging strategies Future research

Definition of hedge ratios Analysis in a market governed by a Markovian contagion model

- Histogram of P&L increments using Δ^{lo} (left) and Δ^{li} (right)
 - Hedging period: [0, 0.09] (one month)
 - $\bar{N} = 10000$ trajectories



- Histogram of P&L increments using Δ^{lo} (left) and Δ^{li} (right)
 - Hedging period: [0, 0.5] (one semester)
 - $\bar{N} = 10000$ trajectories



Products and models under scrutiny Comparison of two hedging strategies Future research

- Standard deviation of P&L increments function of the hedging horizon
 - $\bar{N}=10000$ simulations at each time step
 - $\sigma(\delta P\&L)$ using Δ^{lo} in blue
 - $\sigma(\delta P \& L)$ using Δ^{li} in red



Hedging with individual CDS spreads

- Hedging with individual CDS may perform a better hedge (than hedging with the index)
 - heterogeneous portfolio where some individual spreads are suddenly widening
 - equity tranche very sensitive to idiosyncratic risk
- Obviously, hedging with single name sensitivities is beyond the reach of a pure top model
- Future research: Comparison of hedge performance with individual CDS contracts when hedging strategies are computed
 - using the market standard hedging method (spread-deltas in a base correlation approach): **bottom-up approach**
 - using a pure top model associated with a thinning procedure: top-down approach

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Top-down approach

- In pure top model, the flow of information is only driven by the cumulative loss process $(\mathcal{H}_t = \sigma(L_s, s \leq t))$
 - Given \mathcal{H}_t , we can only forecast the timing of defaults up to time t: ordered default times are \mathcal{H} -stopping times
 - But: lose of information related to the defaulters' identities
- Starting from a top model, Giesecke and Goldberg(2005) propose to recover single name information using a random thinning procedure
- The idea is to allocate a fraction of the loss intensity to each individual name with the constraint that the individual CDS spreads in the portfolio are matched

Set-up

- τ_1, \ldots, τ_n : default time, $N_t = \sum_{i=1}^n \mathbb{1}_{\{\tau_i \leq t\}}$
- Let us define $\mathcal{H} = \{\mathcal{H}_t\}$, where:

$$\mathcal{H}_t = \sigma(N_s, s \le t)$$

• $\tau^{(1)} < \ldots < \tau^{(n)}$: ordered default time

• Let us define by $\mathcal{I} = \{\mathcal{I}_t\}$ the defaulter's identity filtration, where

$$\mathcal{I}_t := \sigma \left(I_{ij} \mid i = 1, \dots, n; \ j = 1, \dots, N_t \right)$$

- $\mathcal{F} = \{\mathcal{F}_t\}$: background filtration that contains the external market information.
- $\mathcal{G} = \{\mathcal{T}_t\}$: largest filtration

$$\mathcal{G}_t = \mathcal{H}_t \vee \mathcal{I}_t \vee \mathcal{F}_t$$

Compensator of ordered default times

- $\tau^{(1)} < \ldots < \tau^{(n)}$ ordered default times are $\mathcal G\text{-stopping time}$
- $\Lambda^{(1)},\ldots,\Lambda^{(n)}$: $\mathcal G$ -compensators of $\tau^{(1)},\ldots,\tau^{(n)}$
- Λ : \mathcal{G} -compensator of N
- Bielecki, Crépey, Jeanblanc(2008):

Proposition

For
$$t\geq 0$$
, $\Lambda_t^{(i)}=\Lambda_{t\wedge au^{(i)}}-\Lambda_{t\wedge au^{(i-1)}}$, $i=1,\ldots,n$

• $\tau^{(1)} < \ldots < \tau^{(n)}$ are $\mathcal H\text{-stopping time}$

Random Thinning

- au_1, \ldots, au_n are \mathcal{G} -stopping times
- $\Lambda_1, \ldots, \Lambda_n$: *G*-compensators of τ_1, \ldots, τ_n
- Λ : \mathcal{G} -compensator of N

$$\Lambda = \sum_{i=1}^{n} \Lambda_i$$

• Giesecke and Goldberg(2005):

Proposition

There exists \mathcal{G} -predictable non-negative processes Z_i , i = 1, ..., n (Z-factors) such that $\sum_{i=1}^{n} Z_i = 1$ and

$$\Lambda_i = \int_0 Z_{i,t} d\Lambda_t, \quad i = 1, \dots, n.$$

Random Thinning

- $\lambda_{i,t}$: *G*-intensities of τ_i , $i = 1, \ldots, n$
- λ_t: G-intensity of N
- $\lambda_{i,t} = Z_{i,t}\lambda_t$, $i = 1, \dots, n$ and $\sum_{i=1}^n Z_i = 1$
- $Z_{i,t}$ is the conditional probability that name *i* is the next defaulter given an imminent default in the interval [t, t + dt]:

$$Z_{i,t} = \sum_{j=1}^{n} \mathbb{P}\left(\tau^{(j)} = \tau_i \mid t < \tau^{(j)} \le t + dt, \,\mathcal{G}_t\right) \mathbf{1}_{\left\{\tau^{(j-1)} < t \le \tau^{(j)}\right\}}$$

• Top-Down matrix: $P(t) = (p_{i,j}(t))_{1 \le i,j \le n}$

$$p_{i,j}(t) = \mathbb{P}\left(\tau^{(j)} = \tau_i \mid t < \tau^{(j)} \le t + dt, \, \mathcal{G}_t\right)$$

• Consistency condition: $\sum_{i=1}^{n} p_{i,j}(t) = 1, j = 1, \dots, n$

Random draws without replacement

- Approach proposed by Halperin and Tomecek (2008)
- TD matrix piecewise constant in t, only change at default times

•
$$t=0$$
, no default $p_{i,j}^0=\mathbb{P}\left(au^{(j)}= au_i
ight)$ (inputs)

- Simulation of $au^{(1)},\ldots, au^{(n)}$ (or N) in the "small filtration", i.e ${\mathcal H}$
- At $t = \tau^{(1)}$ (first jump of N): independent simulation of the defaulter identity $I_1 \in \{1, 2, ..., n\}$ according to the distribution:

$$\mathbb{P}(I_1 = i) = p_{i,1}^0, \ i = 1, \dots, n$$

• Update the TD matrix $p_{i,j}^0 o p_{i,j}^1 = \mathbb{P}\left(au^{(j)} = au_i \mid au^{(1)}, I_1
ight)$

$$p_{i,j}^1 = \begin{cases} 0 & i = I_1, \ j = 1, \dots, n \\ 0 & j = 1, \ i = 1, \dots, n \\ \frac{p_{i,j}^0}{1 - p_{I,j}^0} & i \neq I, \ j = 2, \dots, n \end{cases}$$

• Practical issue: if $N_t = \sum_{i=1}^n \mathbb{1}_{\{\tau_i \leq t\}}$ is Markov with respect to the "small filtration" \mathcal{H} , it is no more the case in the "large filtration" \mathcal{G}

Random draws with replacement

- Pure top model: homogeneous Markovian contagion model (local intensity)
- At the *j*-th jump of N: independent simulation of the defaulter identity $I_j \in \{1, 2, ..., n\}$ according to the distribution:

$$p_{i,j}, \text{ where } \sum_{i=1}^n p_{i,j} = 1$$

- After the draw, I_j is replaced in the pool: TD matrix is not updated
- Denote by $B_{i,j}$, i = 1, ..., n, j = 1, ..., n some independent Bernoulli random variables such that $\mathbb{E}[B_{i,j}] = p_{i,j}$
- We build n individual counting process $N_i(t)$ such that

$$N_{i,t} = \sum_{j=1}^{N_t} B_{i,j}, \ i = 1, \dots, n$$

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Random draws with replacement

- N_i cannot be identified with the "true" usual default process of i (single jump to default)
 - But here $\mathbb{E}[N_T \mid \mathcal{G}_t] = \mathbb{E}[N_T \mid \mathcal{H}_t] = \mathbb{E}[N_T \mid N_t]$
 - Tractable calibration to CDO tranches, individual CDS quotes
 - We hope that individual delta spreads are relevant ...