Dynamic Modeling of Portfolio Credit Risk with Common Shocks

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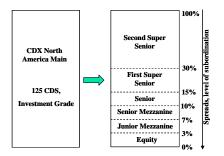




Tom Bielecki, Areski Cousin, Stéphane Crépey and Alexander Herbertsson Dynamic Modeling of Portfolio Credit Risk with Common Shocks

Introduction

Main issue: hedging of portfolio credit derivatives



• Cash-flows driven by the realized path of the aggregate loss process

$$L_{t} = \frac{1}{n} \sum_{i=1}^{n} (1 - R_{i}) H_{t}^{i}$$

where R_i is the recovery rate and H_t^i is the default indicator of obligor i

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Hedging using the one-factor Gaussian copula model?

Advantages:

- **Bottom-up model**: account for dispersion of default risk among names in the portfolio
- Copula construction of default times: Calibration of CDS spreads and CDO tranche quotes can be made using two separate numerical procedures
- Factor model: fast algorithms to compute aggregate loss distribution

Drawbacks:

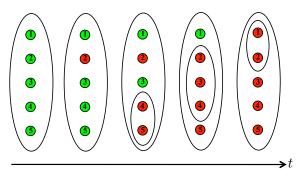
- Static model
- **Base correlation approach** unable to describe consistently the dependence structure of default times

Dynamic model of portfolio credit risk

Simultaneous default model

• Defaults are the consequence of triggering-events affecting simultaneously pre-specified groups of obligors

Example: n = 5 and $\mathcal{Y} = \{\{1\}, \{2\}, \{3\}, \{4\}, \{5\}, \{4, 5\}, \{2, 3, 4\}, \{1, 2\}\}.$



Dynamic model of portfolio credit risk

- $\{1,\ldots,n\}$ set of credit references
- $\mathcal{Y} = \{\{1\}, \dots, \{n\}, I_1, \dots, I_m\}$ pre-specified groups of obligors
- $\lambda_Y = \lambda_Y(t)$ deterministic intensity function of the triggering-event associated with group $Y \in \mathcal{Y}$
- $\mathbf{H}_t = (H_t^1, \dots, H_t^n)$ defined as multivariate continuous-time Markov chain in $\{0, 1\}^n$ such that for $\mathbf{k}, \mathbf{m} \in \{0, 1\}^n$:

$$\mathbb{P}\left(\mathbf{H}_{t+dt} = \mathbf{m} \mid \mathbf{H}_{t} = \mathbf{k}\right) = \sum_{Y \in \mathcal{Y}} \lambda_{Y}(t) \mathbf{1}_{\{\mathbf{k}^{Y} = \mathbf{m}\}} dt$$

where \mathbf{k}^{Y} is obtained from $\mathbf{k} = (k_1, \dots, k_n)$ by replacing the components k_j , $j \in Y$, by number one. ex: $(0, 1, 0, 0)^{\{1, 2, 4\}} = (1, 1, 0, 1)$

• $\mathcal{F}_t = \sigma(\mathbf{H}_u, u \leq t)$ natural filtration of \mathbf{H}

Dynamic model of portfolio credit risk

Example: n = 2, $\mathcal{Y} = \{\{1\}, \{2\}, \{1, 2\}\}$. $\mathbf{H}_t = (H_t^1, H_t^2)$ is a multivariate continuous-time Markov chain with space set $\{(0, 0), (1, 0), (0, 1), (1, 1)\}$ and generator matrix

	(0, 0)	(1,0)	(0,1)	(1,1)	
(0,0)		$\lambda_{\{1\}}$	$\lambda_{\{2\}}$	$\lambda_{\{1,2\}}$	
(1,0)	0	_	0	$\lambda_{\{2\}} + \lambda_{\{1,2\}}$	
(0,1)		0	_	$\lambda_{\{1\}} + \lambda_{\{1,2\}}$	
(1,1)	0	0	0	0)

- $\bullet~$ Obligor 1 defaults with intensity $\lambda_{\{1\}}+\lambda_{\{1,2\}}$ regardless of the state of the pool
- $\bullet~$ Obligor 2 defaults with intensity $\lambda_{\{2\}}+\lambda_{\{1,2\}}$ regardless of the state of the pool
- No contagion effect : Past defaults do not have any effect on intensities of surviving names

General case: Obligor *i* defaults with intensity $\eta_i(t) = \sum_{Y \in \mathcal{Y}} \lambda_Y(t) \mathbf{1}_{\{i \in Y\}}$

$$\mathbb{P}(H_{t+dt}^{i} - H_{t}^{i} = 1 \mid \mathcal{F}_{t}) = \mathbb{P}(H_{t+dt}^{i} - H_{t}^{i} = 1 \mid H_{t}^{i}) = (1 - H_{t}^{i})\eta_{i}(t)dt$$

• Each default indicator $H^i,\,i=1,\ldots,n$ is a Markov process with respect to ${\mathcal F}$

Separate calibration procedure of CDS-s and CDO tranches

For any i = 1, ..., n, the price at time t of a CDS referencing name i (European-type payoff):

$$\mathbb{E}\left[\Phi(H_T^i) \mid \mathcal{F}_t\right] = \mathbb{E}\left[\Phi(H_T^i) \mid H_t^1, \dots, H_t^n\right] = \mathbb{E}\left[\Phi(H_T^i) \mid H_t^i\right]$$

Hedging CDO tranches with single-name CDS

- Derive price dynamics of CDO tranche and single-name CDS-s
- Computation of min-variance hedging strategies in this incomplete market model
- But: price of portfolio loss derivatives solves a large system of Kolmogorov backward equations that is numerically intractable at least for large portfolios (n > 20)

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Common-Shock Model Interpretation

• For any pre-specified group $Y \in \mathcal{Y} = \{\{1\}, \dots, \{n\}, I_1, \dots, I_m\}$, we define

$$\tau_Y = \inf\left\{t \ge 0 \mid \int_0^t \lambda_Y(s) ds > E_Y\right\}$$

where E_Y , $Y \in \mathcal{Y}$ are independent and exponentially distributed random variables with parameter 1.

- τ_Y is the arrival time of shock Y that yields default of non-defaulted names in group Y
- Default time of name $i = 1, \ldots, n$ defined by:

$$\widehat{\tau}_i = \min_{\{Y \in \mathcal{Y}; \, i \in Y\}} \tau_Y$$

Common-Shock Model Interpretation

For all $t_1, \ldots, t_n \ge 0$, the following relation holds

$$\mathbb{P}\left(\widehat{\tau}_1 > t_1, \dots, \widehat{\tau}_n > t_n\right) = \mathbb{P}\left(\tau_1 > t_1, \dots, \tau_n > t_n\right)$$

where $\tau_i:=\inf\left\{t\geq 0\mid H^i_t=1\right\}$ is the default time of name i in the Markovian model

Common-Shock Model Interpretation

Using fast recursion procedure for pricing and hedging CDO tranches

• Thanks to the common-shock model interpretation:

$$L_t = \frac{1}{n} \sum_{i=1}^n (1 - R_i) \mathbf{1}_{\{\tau_i \le t\}} \stackrel{d}{=} \frac{1}{n} \sum_{i=1}^n (1 - R_i) \mathbf{1}_{\{\hat{\tau}_i \le t\}}$$

- where $\mathbf{1}_{\{\hat{\tau}_1 \leq t\}}, \ldots, \mathbf{1}_{\{\hat{\tau}_n \leq t\}}$ are conditionally independent Bernoulli's given common-shock indicators $\mathbf{1}_{\{\tau_{I_1} \leq t\}}, \ldots, \mathbf{1}_{\{\tau_{I_m} \leq t\}}$
- For any state of the Markov process H_t, there exists an "equivalent" common-shock model that matches joint distribution of default times for non-defaulted names.
- Computation of min-variance hedging strategies is also tractable

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Common-Shock Model Interpretation

Two-step calibration procedure

First step:

• CDS spread of name i at time t = 0 can be expressed as a function of survival probabilities $\mathbb{P}(\tau_i > t) = \exp\left(-\int_0^t \eta_i(u) du\right)$ where

$$\eta_i(u) = \lambda_{\{i\}}(u) + \sum_{k=1}^m \lambda_{I_k}(u) \mathbf{1}_{\{i \in I_k\}}$$

 Marginal default intensities η_i, i = 1,..., n, can be calibrated on single-name CDS curves using a standard bootstrap procedure

Second step:

• Common-shock intensities λ_{I_k} , k = 1, ..., m are calibrated on CDO tranche quotes using the recursion algorithm

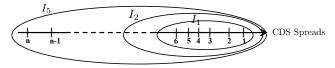
Calibration on CDX index

Data set: 5-year CDX North-America IG index on 20 December 2007

- Quoted spreads at different pillars of the n=125 index constituents
- Quoted spreads of standard tranches [0,3], [3,7], [7,10], [10,15], [15,30]

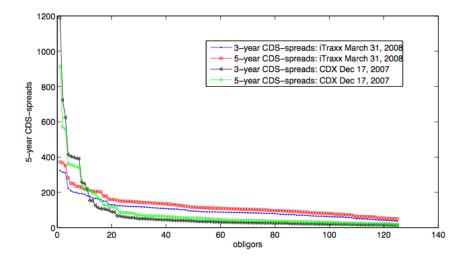
Model specification:

• Names are labelled with respect to decreasing level of spreads



- m = 5 groups $I_1 \subset \cdots \subset I_5$ such that $I_1 = \{1, \dots, 6\}, I_2 = \{1, \dots, 19\}, I_3 = \{1, \dots, 25\}, I_4 = \{1, \dots, 61\}, I_5 = \{1, \dots, 125\}$
- Piecewise-constant intensities $\lambda_{\{1\}}, \ldots, \lambda_{\{125\}}, \lambda_{I_1}, \ldots, \lambda_{I_5}$ with grid points corresponding to CDS pillars
- $\bullet\,$ Homogeneous and constant recovery rates: 40%
- Constant short-term interest rate: 3%

Calibration on CDX index



Calibration results:

Tranche	[0,3]	[3,7]	[7,10]	[10,15]	[15,30]
Model spread in bps	48.0701	254.0000	124.0000	61.0000	38.9390
Market spread in bps	48.0700	254.0000	124.0000	61.0000	41.0000
Abs. Err. in bps	0.0001	0.0000	0.0000	0.0000	2.0610
% Rel. Err.	0.0001	0.0000	0.0000	0.0000	5.0269

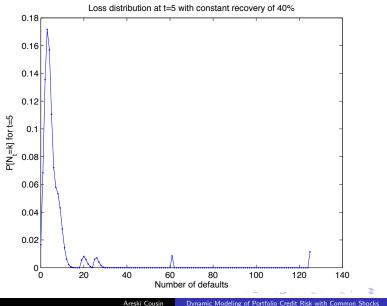
• Names in the set $I_5 \setminus I_4$ are excluded from the calibration constraints (they can only default within the Armageddon shock I_5)

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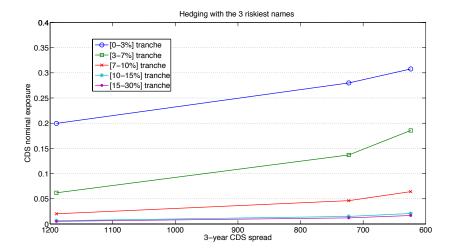
Calibration on CDX index

Implied 5-year loss distribution:

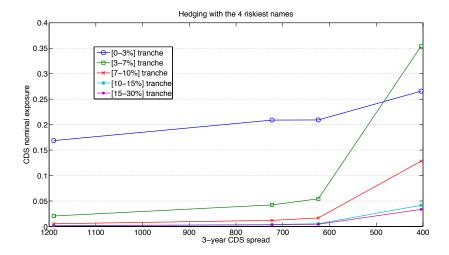


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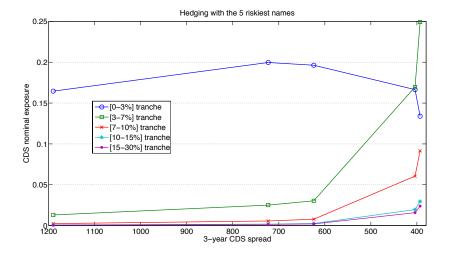
Min-variance hedging strategies



Min-variance hedging strategies

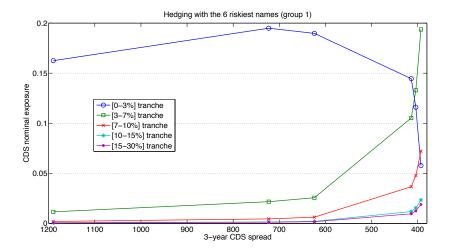


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Min-variance hedging strategies



Thank you for your attention!

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